EC 709: Monotonicity and 2SLS

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2 A Classic Example



2 A Classic Example



Review of Assumptions

A1. Instrument Relevance: $Pr(D = 1|Z = 1) \neq Pr(D = 1|Z = 0)$

- $\bullet\,$ Can be assessed by inferring the coefficient in the first-stage regression: Convention view, $F\geq 10$
- New researches show this is not enough (See Lee et al (2022) and Keane and Neal (2023))
- A2. Exclusion restriction contains two parts:
 - 1. No direct effect on potential outcome: With probability 1, $Y_{d1} = Y_d | Z = 1 = Y_d | Z = 0 = Y_{d0}$ for d = 1, 0
 - Random Assignment: The variable Z is jointly independent of (Y₁₁, Y₁₀, Y₀₁, Y₀₀, D₁, D₀)
 - \Rightarrow Z affect D only by affecting whether the treatment is more likely to be D_1 or D_0
- A3. Monotonicity: With probability 1, the potential treatment response indicators satisfy $D_{1i} \ge D_{0i} \forall i$ or $D_{0i} \ge D_{1i} \forall i$
 - ? How do we test A2 and A3 in practice?

A Necessary Condition on IV validity

• Start with something observable:

$$\begin{aligned} & \Pr(Y = y, D = 1 | Z = 1) - \Pr(Y = y, D = 1 | Z = 0) \\ &= (\Pr(Y_1 = y, D_1 = 1, D_0 = 0 | Z = 1) + \Pr(Y_1 = y, D_1 = 1, D_0 = 1 | Z = 1)) \\ &- (\Pr(Y_1 = y, D_1 = 1, D_0 = 1 | Z = 0) + \Pr(Y_1 = y, D_1 = 0, D_0 = 1 | Z = 0)) \\ & (By \text{ Random assignment of } Z \text{ to potential outcomes:}) \\ &= \Pr(Y_1 = y, D_1 = 1, D_0 = 0 | Z = 1) - \Pr(Y_1 = y, D_1 = 0, D_0 = 1 | Z = 0) \\ &= \Pr(Y_1 = y, D_1 = 1, D_0 = 0) - \Pr(Y_1 = y, D_1 = 0, D_0 = 1) \end{aligned}$$

• Further assuming $D_{1i} \ge D_{0i} \forall i$ and by the similar procedure:

 $Pr(Y = y, D = 1 | Z = 1) - Pr(Y = y, D = 1 | Z = 0) = Pr(Y_1 = y, D_1 > D_0)$ $Pr(Y = y, D = 0 | Z = 0) - Pr(Y = y, D = 0 | Z = 1) = Pr(Y_0 = y, D_1 > D_0)$

 \Rightarrow Testable implication: nonnegative difference in densities

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A Test For Instrument Validity (Kitagawa, 2015)

p(y,d) = Pr(Y = y, D = d | Z = 1), q(y,d) = Pr(Y = y, D = d | Z = 0)

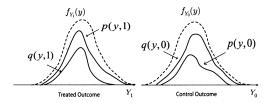


Figure 1: IV validity cannot be refuted

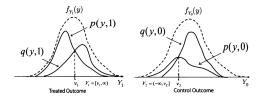


Figure 2: Can refute at least one of the IV validity assumptions

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3 General 2SLS and LATE

Card (1995)

- Data: 24-year-old men from the 1976 interview of the NLSYM, N = 3,010
 - National Longitudinal Survey of Young Men (NLSYM): sampled men aged 14–24 in 1966 and continued with follow-up surveys through 1981
- Question: estimate the returns to education
 - The outcome Y: log hourly wage; The treatment D: indicates whether one graduated from a four-year college
 - Endogenity: Omitted ability measure
- The instrument Z: a binary indicator for the presence of an accredited four-year college in the local labor market when the respondent was 14 years old
 - "the distance to the nearest college" as an instrument for educational attainment
 - First Stage: Live in an area close to college makes students more likely to attend college
 - the presence of a nearby college reduces the cost of college education by allowing students to live at home

Testing For Card (1995)

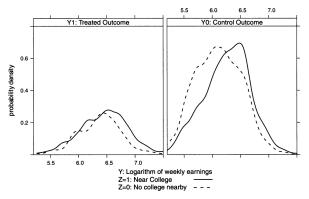


FIGURE 4.—Kernel density estimates: proximity to college data. The Gaussian kernel with bandwidth 0.08 is used.

- Either Exclusion restriction or Monotonicity failed
- $\Rightarrow\,$ Need to diagnose by economics intuition and incorporate more general methods

Problem of the Exclusion restriction

- Implication of A2: Student's unobservable ability is presumably independent of students' residence during their teenage years
 - Do not control for any demographic covariates
 - ⇒ Raises a concern regarding the violation of the random assignment assumption
 - E.g., Urban areas are more likely to have colleges and higher wage levels compared to the Rural areas
- After adding some additional covariates, Kitagawa (2015) failed to reject the validity of Card (1995)'s instrument:
 - Five binary variables: whether Black, whether lived in a metropolitan area (SMSA) in 1966 and 1976, and whether lived in the South in 1966 and 1976
- In Card's main results (1995, Table 3A, column (5)), he indeed emphasized the importance of adding additional controls
- Based on a survey by Blandhol et al (2022) , 81% of papers using IV included at least one covariate X

- Compilers: Students grew up in relatively low-income families and who were not able to go to college without living with their parents
- Frolich and Sperlich (2019): Might not be the only direction
 - Some students may be encouraged to attend college if the nearest college is far away, as this gives them an excuse to move out of the parental home
 - ← Defier: "defy" their instrument assignment for any reason
- Monotonicity assumption is likely to fail

- Weak Monotonicity (WM): There exists a partition of the covariate space such that P[D₁ ≥ D₀|X] = 1 a.s. on one subset and P[D₁ ≤ D₀|X] = 1 a.s. on its complement
 - Defiers we mentioned above are unlikely to affect students with binding financial constraints
 - It is conceivable that college proximity never discourages poor students from attending college and never encourages rich students to do so
 - ⇒ Consistent with the Weak Monotonicity with the partition on the income level of the households
 - $\Leftarrow\,$ Direction itself is allowed to be different for the two groups
- Obviously true that Assumption WM is weaker than Monotonicity
- Still restrictive to assume that all rich students and all poor students are affected by college proximity in the same direction (if at all)

- In practice, we often need to **add some covariates** and **assume weak monotonicity** to achieve a valid IV
- Are we still estimating LATE in our 2SLS with these further generalizations?
- \Rightarrow I will offer a quick review of what is trending in the IV literature
 - ! Make sure the covariates you add are exogenous!
- Glynn and Rueda (2017) called that post-instrument bias if covariates are itself endogenous
- \Rightarrow OLS with an omitted variable will often have less bias than IV with the post-instrument covariate

2 A Classic Example

General 2SLS and LATE

Covariates + Monotonicity

• X: a vector of control variables including a constant; $P[D_1 \ge D_0] = 1$

•
$$\beta_{iv} = \frac{E[Y\tilde{Z}]}{E[T\tilde{Z}]}$$
, where $\tilde{Z} = Z - L[Z|X]$

- \tilde{Z} : residuals from a regression of Z on X
- L[Z|X] = X^TE[XX^T]⁻¹E[XZ]: instrument propensity score when Z is binary
- Blandhol et al (2022) shows:

 $\beta_{iv} = E[\omega(cp, X)\tau(cp, X)] + E[\omega(at, X)\tau(at, X)] + E[\omega(nt, X)\tau(nt, X)]$

- τ(T, X): conditional average treatment effects for group T; ω(T, X): weights on group T
- Whenever $L[Z|X] \neq E[Z|X]$, the IV estimand incorporates negatively weighted treatment effects for some groups

- Now given that L[Z|X] = E[Z|X], when we only assumed the Weak Monotonicity :
- Słoczynski (2022) shows there would also be negative weights
- Reduced-form and first-stage regressions implicitly restrict the effects of the instrument to be homogeneous and are thus possibly misspecified
 - Even if all weights are positive, the IV estimand in the just identified specification is not interpretable as the unconditional LATE
 - when almost no individuals are encouraged to get treated, the IV estimand is similar to the local average treatment effect on the treated
 - The opposite of what we want if our goal is to recover the unconditional LATE parameter
 - Słoczynski (2020) finds the similar phenomena for OLS as well

- Both papers use the Card(1995) for empirical application
- Use 2SLS, college attendance yields earnings gains of about 60 log points,
 - Outside the range of estimates in the recent literature
 - - Corrected estimates indicate that attending college causes earnings to be roughly 20% higher

- 1. Misspecification of the model for the instrument propensity score could lead to a large bias
- 2. 2SLS implicitly restricts the effects of the instrument to be homogeneous
- 3. Not desired weight for unconditional LATE
- $\Rightarrow~$ Motivate the use of matching techniques

Thank You!